

The Unscented Particle Filter

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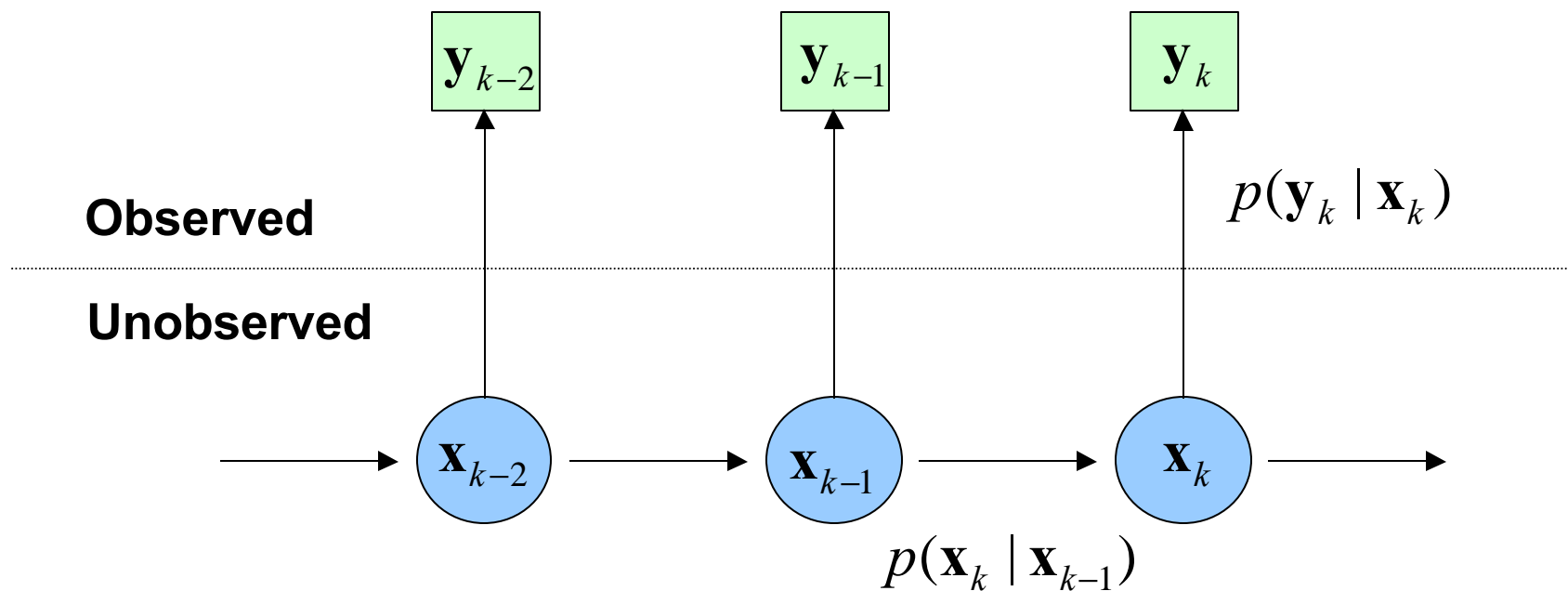
(OGI)

Outline

- **Optimal Estimation & Filtering**
- **Optimal Recursive Bayesian Solution**
- **Practical Solutions**
 - ◆ Gaussian approximations (EKF, UKF)
 - ◆ Sequential Monte Carlo methods (Particle Filters)
- **The Unscented Particle Filter**
 - ◆ The Unscented Transformation and UKF
 - ◆ Applications of UT/UKF to Particle Filters
- **Experimental Results**
- **Conclusions**

Filtering

- General problem statement



- ◆ Filtering is the problem of sequentially estimating the states (parameters or hidden variables) of a system as a set of observations become available on-line.

Filtering

- **Solution of sequential estimation problem given by**

- ◆ Posterior density :

$$p(\mathbf{X}_k | \mathbf{Y}_k)$$

$$\mathbf{X}_k = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_k\}$$

$$\mathbf{Y}_k = \{\mathbf{y}_1, \mathbf{y}_2, \dots, \mathbf{y}_k\}$$

- ◆ By recursively computing a marginal of the posterior, the *filtering density*,

$$p(\mathbf{x}_k | \mathbf{Y}_k)$$

one need not keep track of the complete history of the states.

Filtering

- **Given the filtering density, a number of estimates of the system state can be calculated:**

- ◆ Mean (optimal MMSE estimate of state)

$$\hat{\mathbf{x}}_k = E[\mathbf{x}_k | \mathbf{Y}_k] = \int \mathbf{x}_k p(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k$$

- ◆ Mode
- ◆ Median
- ◆ Confidence intervals
- ◆ Kurtosis, etc.

State Space Formulation of System

- General discrete-time nonlinear, non-Gaussian dynamic system

The diagram shows two equations on a light green background. The first equation is $\mathbf{x}_k = \mathbf{f}(\mathbf{x}_{k-1}, \mathbf{u}_{k-1}, \mathbf{v}_{k-1})$. A red arrow points from the word "state" to \mathbf{x}_k . Another red arrow points from "process noise" to \mathbf{v}_{k-1} . The second equation is $\mathbf{y}_k = \mathbf{h}(\mathbf{x}_k, \mathbf{u}_k, \mathbf{n}_k)$. A red arrow points from "noisy observation" to \mathbf{y}_k . Another red arrow points from "known input" to \mathbf{u}_k . A third red arrow points from "measurement noise" to \mathbf{n}_k .

$$\mathbf{x}_k = \mathbf{f}(\mathbf{x}_{k-1}, \mathbf{u}_{k-1}, \mathbf{v}_{k-1})$$
$$\mathbf{y}_k = \mathbf{h}(\mathbf{x}_k, \mathbf{u}_k, \mathbf{n}_k)$$

- ◆ Assumptions :

- 1) States follow a first order Markov process

$$p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{x}_{k-2}, \dots, \mathbf{x}_0) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$$

- 2) Observations independent given the states

$$p(\mathbf{y}_k | \mathbf{x}_k, A) = p(\mathbf{y}_k | \mathbf{x}_k)$$

Recursive Bayesian Estimation

- Given this state space model, how do we recursively estimate the filtering density ?

$$\begin{aligned} p(\mathbf{x}_k | \mathbf{Y}_k) &= \frac{p(\mathbf{Y}_k | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{Y}_k)} \\ &= \frac{p(\mathbf{y}_k, \mathbf{Y}_{k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{y}_k, \mathbf{Y}_{k-1})} \\ &= \frac{p(\mathbf{y}_k | \mathbf{Y}_{k-1}, \mathbf{x}_k) p(\mathbf{Y}_{k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{y}_k | \mathbf{Y}_{k-1}) p(\mathbf{Y}_{k-1})} \\ &= \frac{p(\mathbf{y}_k | \mathbf{Y}_{k-1}, \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Y}_{k-1}) p(\mathbf{Y}_{k-1}) p(\mathbf{x}_k)}{p(\mathbf{y}_k | \mathbf{Y}_{k-1}) p(\mathbf{Y}_{k-1}) p(\mathbf{x}_k)} \\ &= \frac{p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Y}_{k-1})}{p(\mathbf{y}_k | \mathbf{Y}_{k-1})} \end{aligned}$$

Recursive Bayesian Estimation

likelihood prior

$$p(\mathbf{x}_k | \mathbf{Y}_k) = \frac{p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Y}_{k-1})}{p(\mathbf{y}_k | \mathbf{Y}_{k-1})}$$

posterior evidence

The diagram shows the equation for the posterior probability $p(\mathbf{x}_k | \mathbf{Y}_k)$. Red arrows point from the labels 'likelihood', 'prior', 'posterior', and 'evidence' to their respective parts in the equation. 'likelihood' points to $p(\mathbf{y}_k | \mathbf{x}_k)$, 'prior' points to $p(\mathbf{x}_k | \mathbf{Y}_{k-1})$, 'posterior' points to the entire fraction, and 'evidence' points to the denominator $p(\mathbf{y}_k | \mathbf{Y}_{k-1})$.

- ◆ **Prior :** $p(\mathbf{x}_k | \mathbf{Y}_{k-1}) = \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{Y}_{k-1}) d\mathbf{x}_{k-1}$

transition density given by *process model*

(Propagation of past state into future before new observation is made.)

- ◆ **Likelihood :** defined in terms of *observation model*

- ◆ **Evidence :** $p(\mathbf{y}_k | \mathbf{Y}_{k-1}) = \int p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Y}_{k-1}) d\mathbf{x}_k$

Practical Solutions

- **Gaussian Approximations**
- **Perfect Monte Carlo Simulation**
- **Sequential Monte Carlo Methods : “Particle Filters”**
 - ◆ Bayesian Importance Sampling
 - ◆ Sampling-importance resampling (SIR)

Gaussian Approximations

- Most common approach.
- Assume all RV statistics are *Gaussian*.
- Optimal recursive MMSE estimate is then given by

$$\hat{\mathbf{x}}_k = (\text{prediction of } \mathbf{x}_k) + \mathbf{k}_k [(\text{observation of } \mathbf{y}_k) - (\text{prediction of } \mathbf{y}_k)]$$

- Different implementations :
 - ◆ *Extended Kalman Filter* (EKF) : optimal quantities approximated via first order Taylor series expansion (linearization) of process and measurement models.
 - ◆ *Unscented Kalman Filter* (UKF) : optimal quantities calculated using the Unscented Transformation (accurate to second order for any nonlinearity). Drastic improvement over EKF [Wan, van der Merwe, Nelson 2000].
- Problem : Gaussian approximation breaks down for most nonlinear real-world applications (multi-modal distributions, non-Gaussian noise sources, etc.)

Perfect Monte Carlo Simulation

- Allow for a complete representation of the posterior distribution.
- Map intractable integrals of optimal Bayesian solution to tractable discrete sums of weighted samples drawn from the posterior distribution.

$$\hat{p}(\mathbf{x}_k | \mathbf{Y}_k) = \frac{1}{N} \sum_{i=1}^N \mathbf{d}(\mathbf{x}_k - \mathbf{x}_k^{(i)})$$

$$\mathbf{x}_k^{(i)} \xleftarrow{\text{I.I.D.}} p(\mathbf{x}_k | \mathbf{Y}_k)$$

- So, any estimate of the form

$$E[f(\mathbf{x}_k)] = \int f(\mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k$$

may be approximated by :

$$E[f(\mathbf{x}_k)] \approx \frac{1}{N} \sum_{i=1}^N f(\mathbf{x}_k^{(i)})$$

Particle Filters

- **Bayesian Importance Sampling**

- ◆ It is often impossible to sample directly from the true posterior density.
- ◆ However, we can rather sample from a known, easy-to-sample, *proposal* distribution,

$$q(\mathbf{x}_k | \mathbf{Y}_k)$$

and make use of the following substitution

$$\begin{aligned} E[f(\mathbf{x}_k)] &= \int f(\mathbf{x}_k) \frac{p(\mathbf{x}_k | \mathbf{Y}_k)}{q(\mathbf{x}_k | \mathbf{Y}_k)} q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k \\ &= \int f(\mathbf{x}_k) \frac{p(\mathbf{Y}_k | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{Y}_k) q(\mathbf{x}_k | \mathbf{Y}_k)} q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k \\ &= \int f(\mathbf{x}_k) \frac{w_k(\mathbf{x}_k)}{p(\mathbf{Y}_k)} q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k \\ w_k(\mathbf{x}_k) &= \frac{p(\mathbf{Y}_k | \mathbf{x}_k) p(\mathbf{x}_k)}{q(\mathbf{x}_k | \mathbf{Y}_k)} \end{aligned}$$

Particle Filters

$$\begin{aligned} E[f(\mathbf{x}_k)] &= \frac{1}{p(\mathbf{Y}_k)} \int f(\mathbf{x}_k) w_k(\mathbf{x}_k) q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k \\ &= \frac{\int f(\mathbf{x}_k) w_k(\mathbf{x}_k) q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k}{\int p(\mathbf{Y}_k | \mathbf{x}_k) p(\mathbf{x}_k) \frac{q(\mathbf{x}_k | \mathbf{Y}_k)}{q(\mathbf{x}_k | \mathbf{Y}_k)} d\mathbf{x}_k} \\ &= \frac{\int f(\mathbf{x}_k) w_k(\mathbf{x}_k) q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k}{\int w_k(\mathbf{x}_k) q(\mathbf{x}_k | \mathbf{Y}_k) d\mathbf{x}_k} \\ &= \frac{E_{q(\mathbf{x}_k | \mathbf{Y}_k)} [w_k(\mathbf{x}_k) f(\mathbf{x}_k)]}{E_{q(\mathbf{x}_k | \mathbf{Y}_k)} [w_k(\mathbf{x}_k)]} \end{aligned}$$

Particle Filters

- ◆ So, by drawing samples from $q(\mathbf{x}_k | \mathbf{Y}_k)$, we can approximate expectations of interest by the following:

$$\begin{aligned} E[f(\mathbf{x}_k)] &\approx \frac{\frac{1}{N} \sum_{i=1}^N w_k(\mathbf{x}_k^{(i)}) f(\mathbf{x}_k^{(i)})}{\frac{1}{N} \sum_{i=1}^N w_k(\mathbf{x}_k^{(i)})} \\ &\approx \sum_{i=1}^N \tilde{w}_k(\mathbf{x}_k^{(i)}) f(\mathbf{x}_k^{(i)}) \end{aligned}$$

- ◆ Where the normalized importance weights are given by

$$\tilde{w}_k(\mathbf{x}_k^{(i)}) = \frac{w_k(\mathbf{x}_k^{(i)})}{\sum_{j=1}^N w_k(\mathbf{x}_k^{(j)})}$$

Particle Filters

- ◆ Using the state space assumptions (1st order Markov / observational independence given state), the importance weights can be estimated recursively by [proof in De Freitas (2000)]

$$w_k = w_{k-1} \frac{p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{x}_{k-1})}{q(\mathbf{x}_k | \mathbf{X}_{k-1}, \mathbf{Y}_k)}$$

- ◆ Problem with SIS is that the variance of the importance weights increase stochastically over time [Kong et al. (1994), Doucet et al. (1999)]
- ◆ To solve this, we need to resample the particles
 - keep / multiply particles with high importance weights
 - discard particles with low importance weights
- ◆ *Sampling-importance Resampling (SIR)*

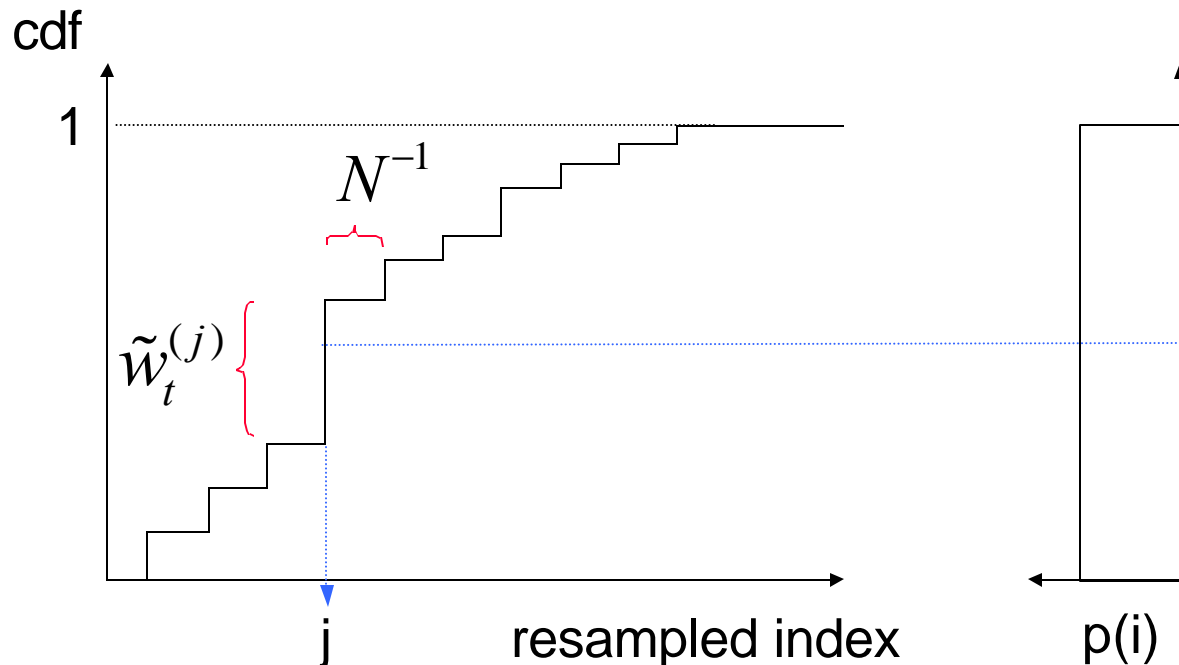
Particle Filters

- **Sampling-importance Resampling**

- ◆ Maps the N unequally weighted particles into a new set of N equally weighted samples.

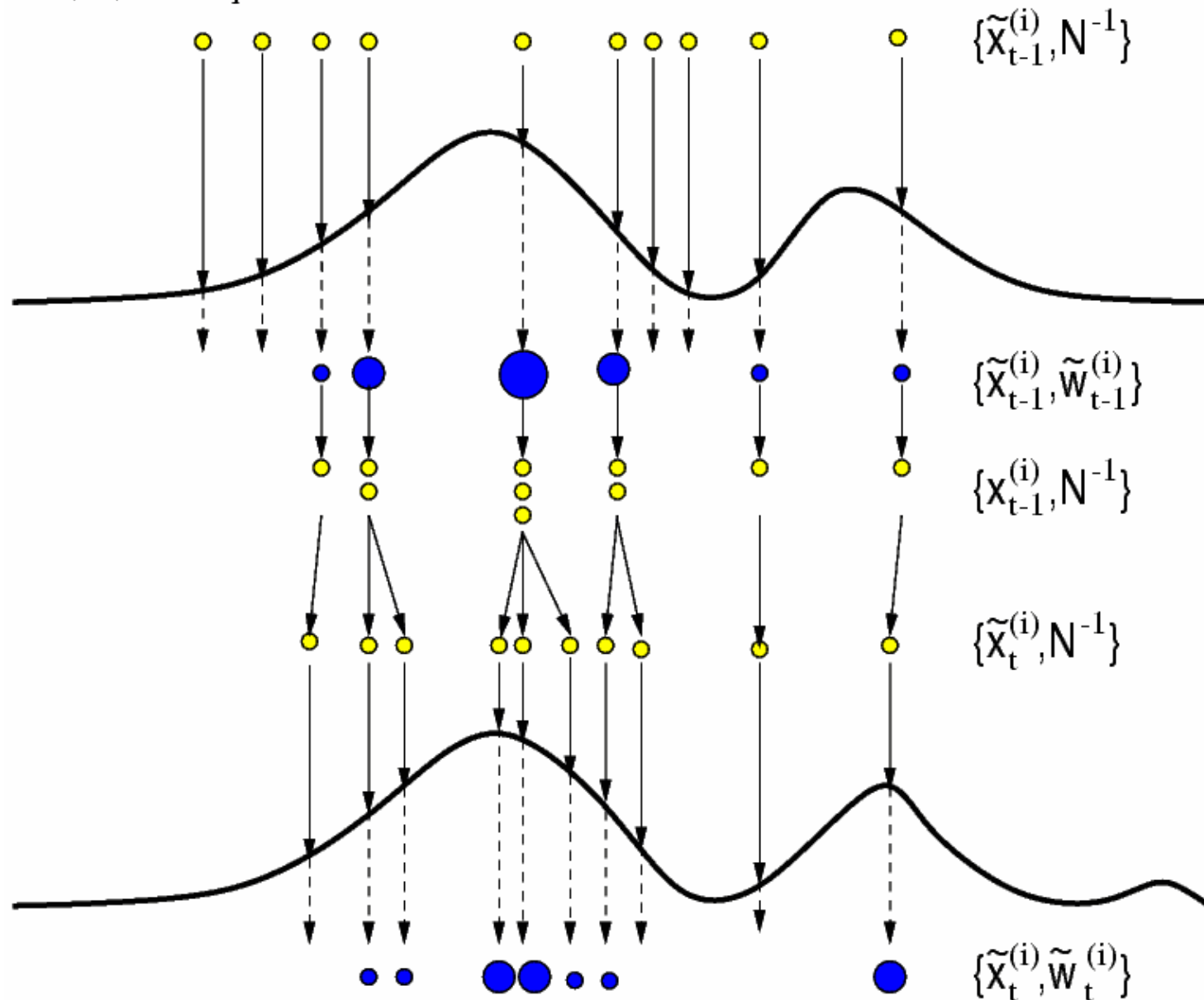
$$\{\mathbf{x}_k^{(i)}, \tilde{w}_k^{(i)}\} \rightarrow \{\mathbf{x}_k^{(j)}, N^{-1}\}$$

- ◆ Method proposed by Gordon, Salmond & Smith (1993) and proven mathematically by Gordon (1994).



Particle Filters

$i=1, \dots, N=10$ particles



Particle Filters

- **Choice of Proposal Distribution**

$$w_k = w_{k-1} \frac{p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{x}_{k-1})}{q(\mathbf{x}_k | \mathbf{X}_{k-1}, \mathbf{Y}_k)}$$

→ critical design issue for successful particle filter

- samples/particles are drawn from this distribution
- used to evaluate importance weights

- ◆ **Requirements**

- 1) Support of proposal distribution must include support of true posterior distribution, i.e. *heavy-tailed* distributions are preferable.
- 2) Must include most recent observations.

Particle Filters

- ◆ Most popular choice of proposal distribution does not satisfy these requirements though:

$$q(\mathbf{x}_k | \mathbf{X}_{k-1}, \mathbf{Y}_k) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$$

[Isard and Blake 96, Kitagawa 96, Gordon et al. 93, Beadle and Djuric 97, Avitzour 95]

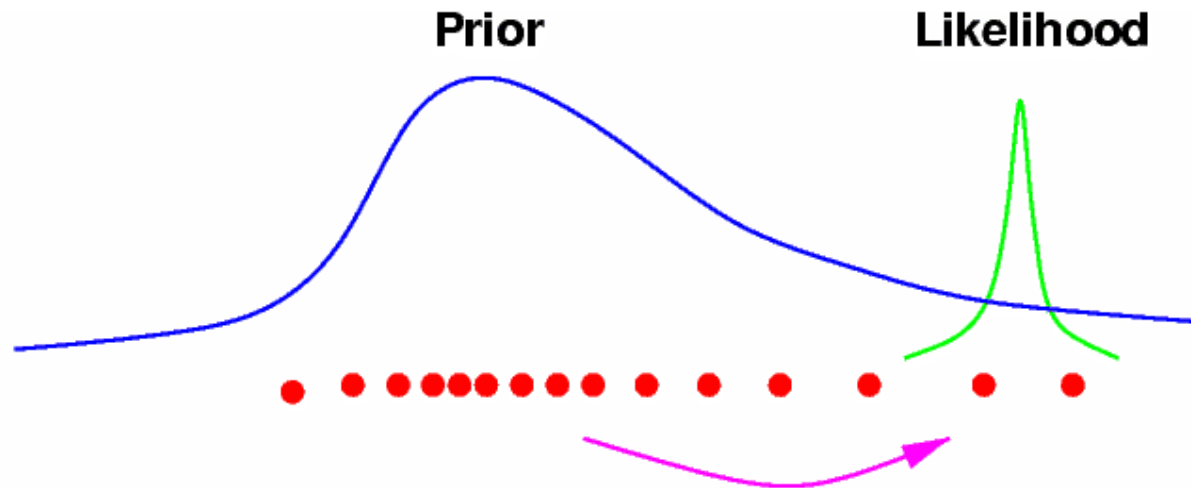
- ◆ Easy to implement :

$$\begin{aligned} w_k &= w_{k-1} \frac{p(\mathbf{y}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{x}_{k-1})}{p(\mathbf{x}_k | \mathbf{x}_{k-1})} \\ &= w_{k-1} p(\mathbf{y}_k | \mathbf{x}_k) \end{aligned}$$

- ◆ Does not incorporate most recent observation though !

Improving Particle Filters

- Incorporate New Observations into Proposal



- ◆ Use Gaussian approximation (i.e. Kalman filter) to generate proposal by combining new observation with prior

$$\begin{aligned} q(\mathbf{x}_k | \mathbf{X}_{k-1}, \mathbf{Y}_k) &= p_G(\mathbf{x}_k | \mathbf{X}_{k-1}, \mathbf{Y}_k) \\ &= \mathcal{N}(\hat{\mathbf{x}}_k, \text{cov}[\mathbf{x}_k]) \end{aligned}$$

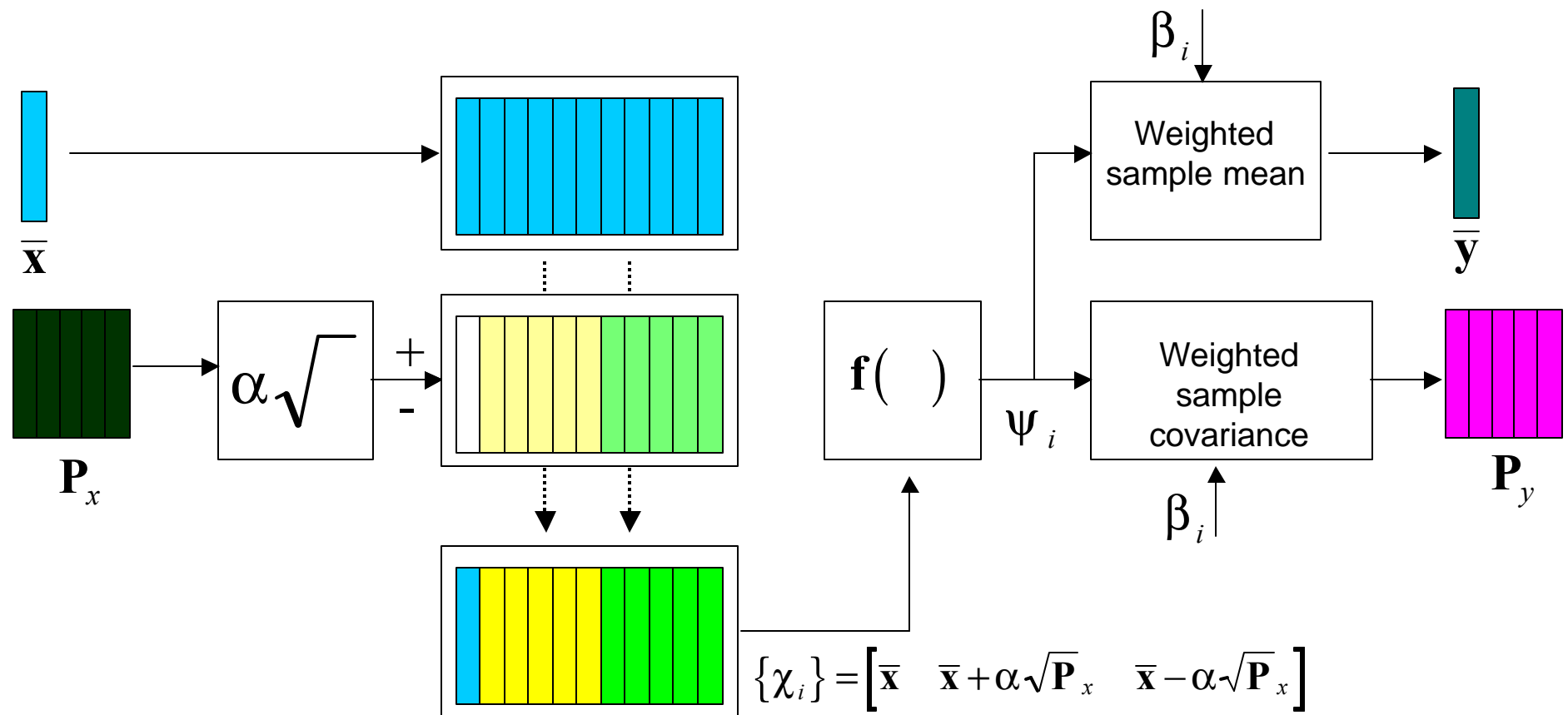
Improving Particle Filters

- ***Extended Kalman Filter Proposal Generation***
 - ◆ De Freitas (1998), Doucet (1998), Pitt & Shephard (1999).
 - ◆ Greatly improved performance compared to standard particle filter in problems with very accurate measurements, i.e. likelihood very peaked in comparison to prior.
 - ◆ In highly nonlinear problems, the EKF tends to be very inaccurate and *underestimates* the true covariance of the state. This violates the distribution support requirement for the proposal distribution and can lead to poor performance and filter divergence.
- We propose the use of the ***Unscented Kalman Filter*** for proposal generation to address these problems !

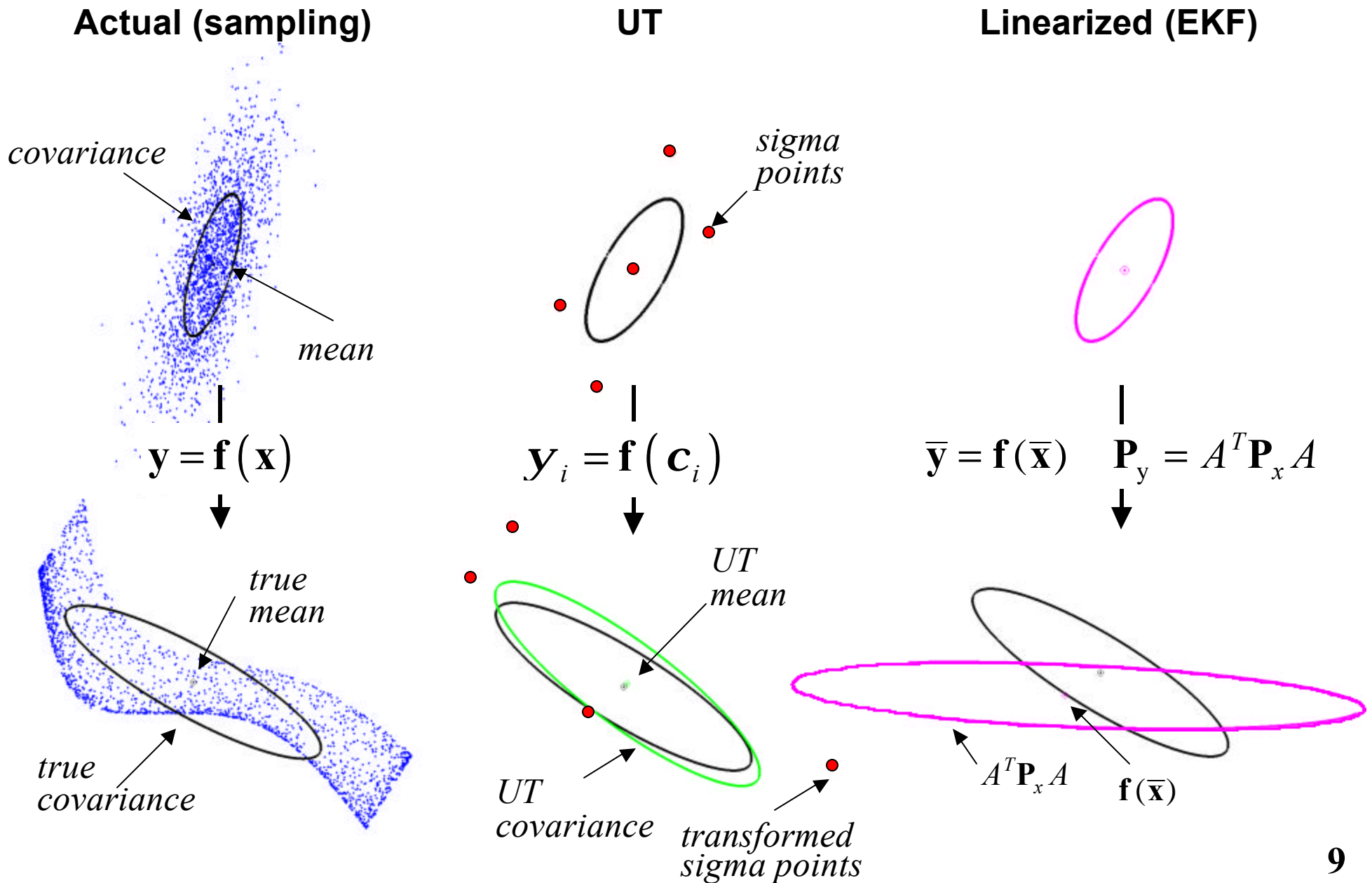
Improving Particle Filters

- ***Unscented Kalman Filter Proposal Generation***
 - ◆ UKF is a recursive MMSE estimator based on the Unscented Transformation (UT).
 - ◆ UT : Method for calculating the statistics of a RV that undergoes a nonlinear transformation (Julier and Uhlmann 1997)
 - ◆ UT/UKF : - accurate to 3rd order for Gaussians
 - higher order errors scaled by choice of transform parameters.
 - ◆ More accurate estimates than EKF (Wan, van der Merwe, Nelson 2000)
 - ◆ Have some control over higher order moments, i.e. kurtosis, etc. → heavy tailed distributions !

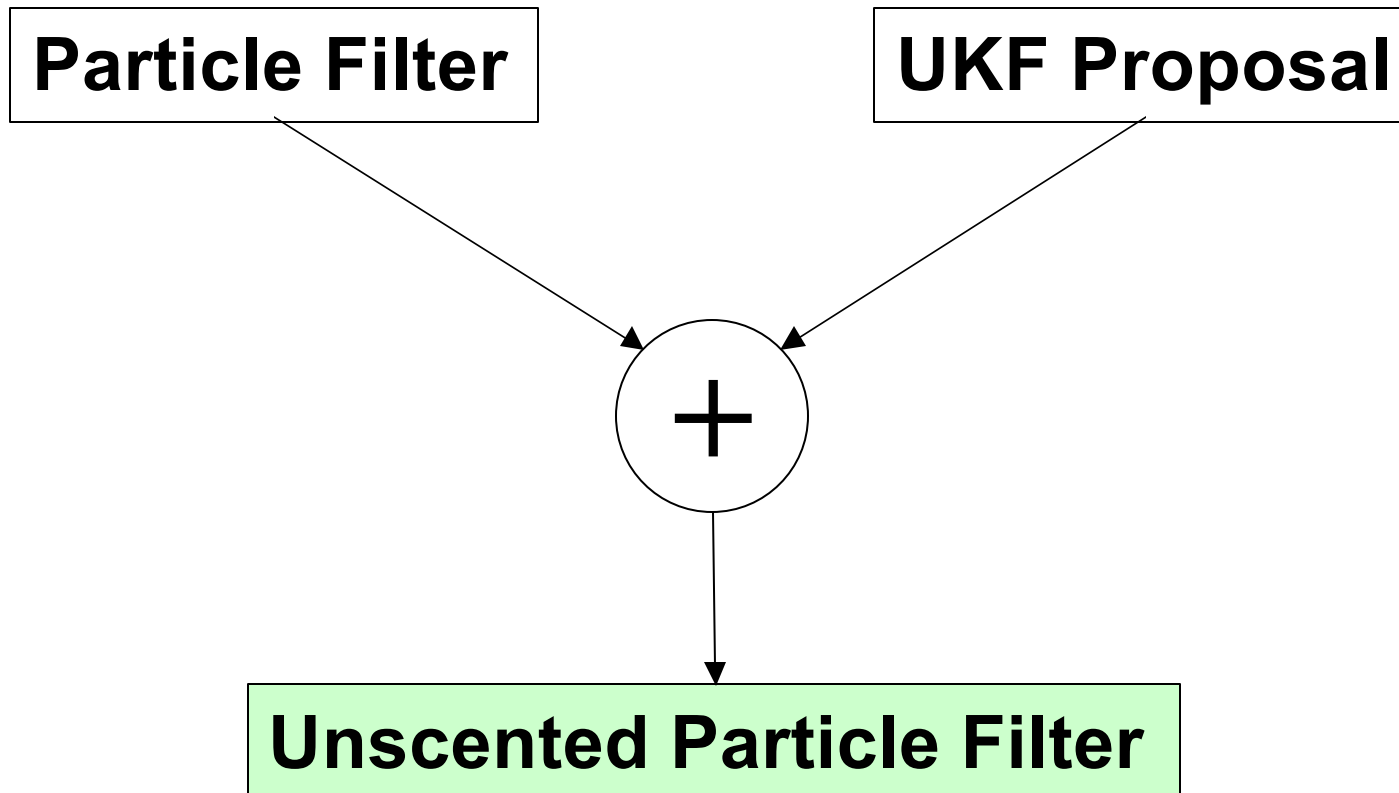
Unscented Transformation



The Unscented Transformation



Unscented Particle Filter



Experimental Results

- Synthetic Experiment

- ◆ Time-series

- process model :

$$x_{k+1} = 1 + \sin(\mathbf{w}p k) + \mathbf{f}x_k + v_k$$

process noise (Gamma) 

- nonstationary observation model :

$$y_k = \begin{cases} \mathbf{f}x_k^2 + n_k & k \leq 30 \\ \mathbf{f}x_k - 2 + n_k & k > 30 \end{cases}$$

 measurement noise (Gaussian)

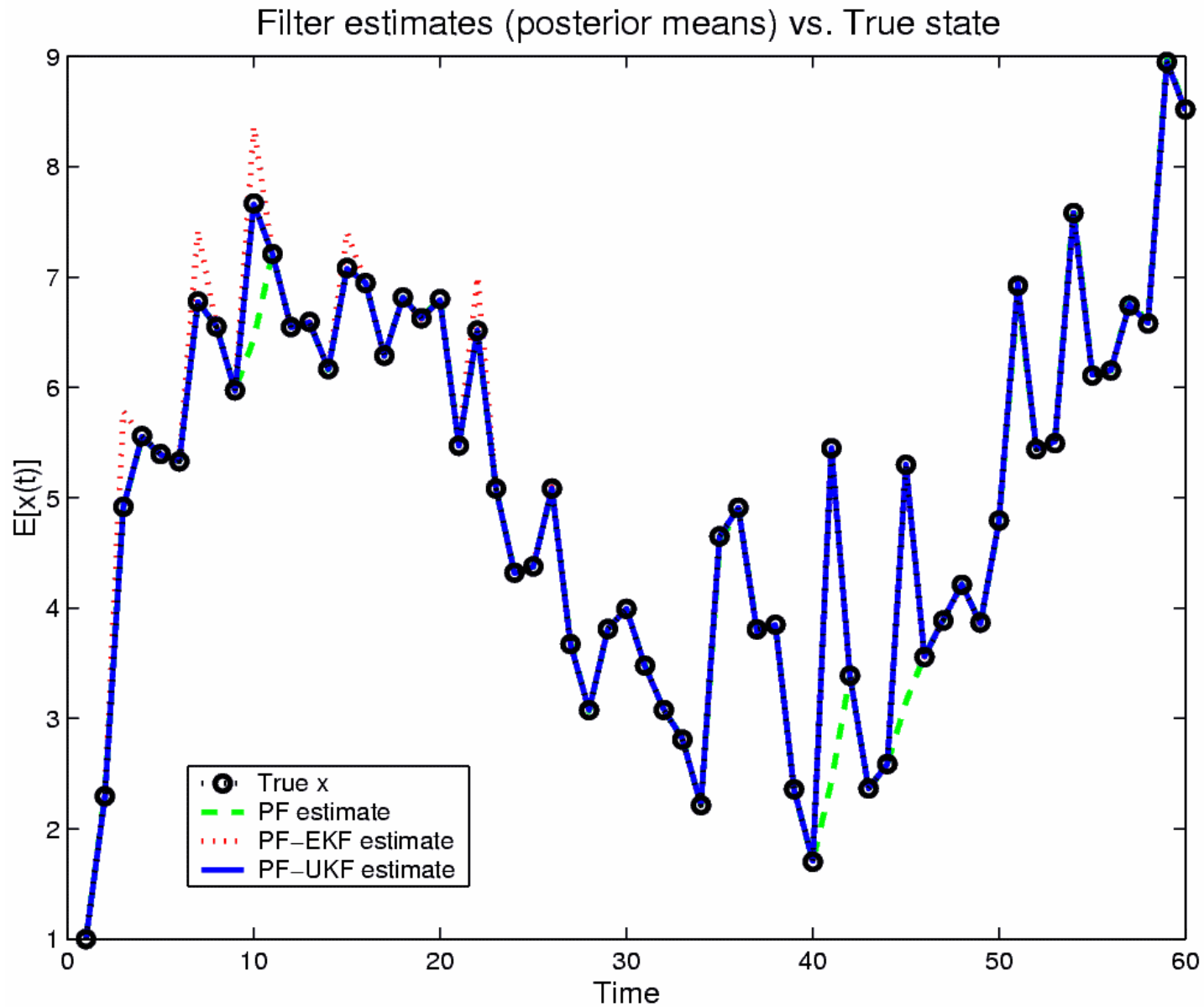
Experimental Results

- **Synthetic Experiment : (100 independent runs)**

Filter	MSE	
	mean	variance
Extended Kalman Filter (EKF)	0.374	0.015
Unscented Kalman Filter (UKF)	0.280	0.012
Particle Filter : generic	0.424	0.053
Particle Filter : EKF proposal	0.310	0.016
<i>Unscented Particle Filter</i>	<i>0.070</i>	<i>0.006</i>

Experimental Results

- Synthetic Experiment



Experimental Results

● Pricing Financial Options

- ◆ Options : financial derivative that gives the holder the right (but not obligation) to do something in the future.
 - Call option : - allow holder to *buy* an underlying cash product
 - at a *specified future date* (“maturity time”)
 - for a *predetermined price* (“strike price”)
 - Put option : - allow holder to *sell* an underlying cash product
- ◆ *Black Scholes* partial differential equation
 - Main industry standard for pricing options

$$\frac{\partial f}{\partial t} + rS \frac{\partial f}{\partial S} + \frac{1}{2} \sigma^2 S^2 \frac{\partial^2 f}{\partial S^2} = rf$$

Diagram illustrating the Black-Scholes partial differential equation with labels:

- risk-free interest rate (points to r)
- value of underlying cash product (points to S)
- volatility of cash product (points to σ)
- option value (points to f)

Experimental Results

- **Pricing Financial Options**

- ◆ Black & Scholes (1973) derived the following pricing solution:

$$C = S\mathcal{N}_c(d_1) - Xe^{-rt_m}\mathcal{N}_c(d_2)$$

$$P = -S\mathcal{N}_c(-d_1) + Xe^{-rt_m}\mathcal{N}_c(-d_2)$$

$$d_1 = \frac{\ln(S/X) + (r + \frac{1}{2}\mathbf{S}^2)t_m}{\mathbf{S}\sqrt{t_m}}$$

$$d_2 = d_1 - \mathbf{S}\sqrt{t_m}$$

$\mathcal{N}_c(\cdot)$ = cumulative normal distribution

Experimental Results

- **Pricing Financial Options**

- ◆ State-space representation to model system for particle filters

- Hidden states : r , S
- Output observations: C , P
- Known control signals: t_m , S

- ◆ Estimate call and put prices over a 204 day period on the FTSE-100 index.

- Performance : normalized square error for one-step-ahead predictions

$$NSE = \sqrt{\sum_k (y_k - \hat{y}_k)^2}$$

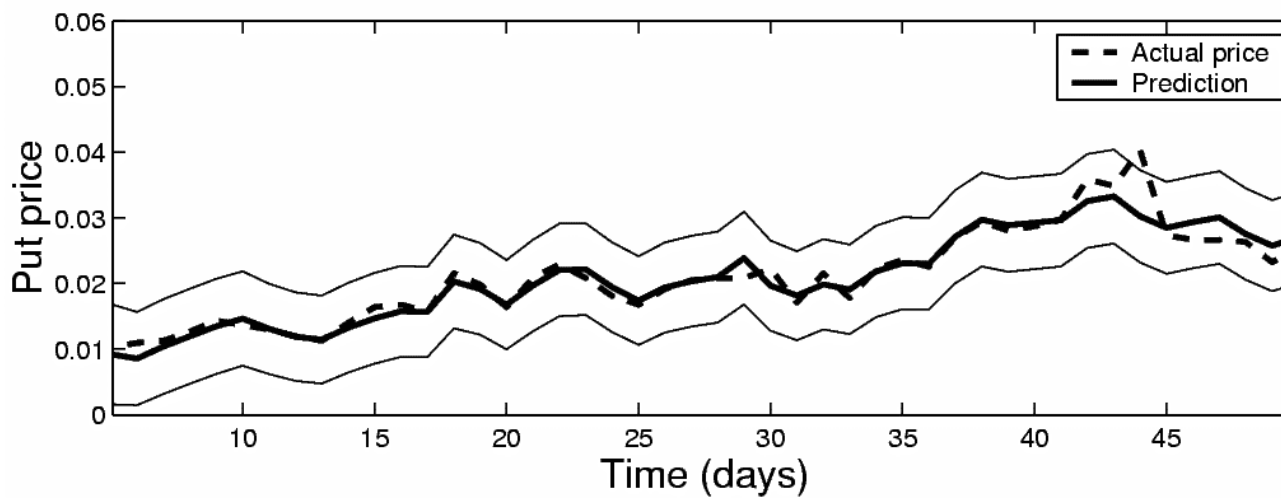
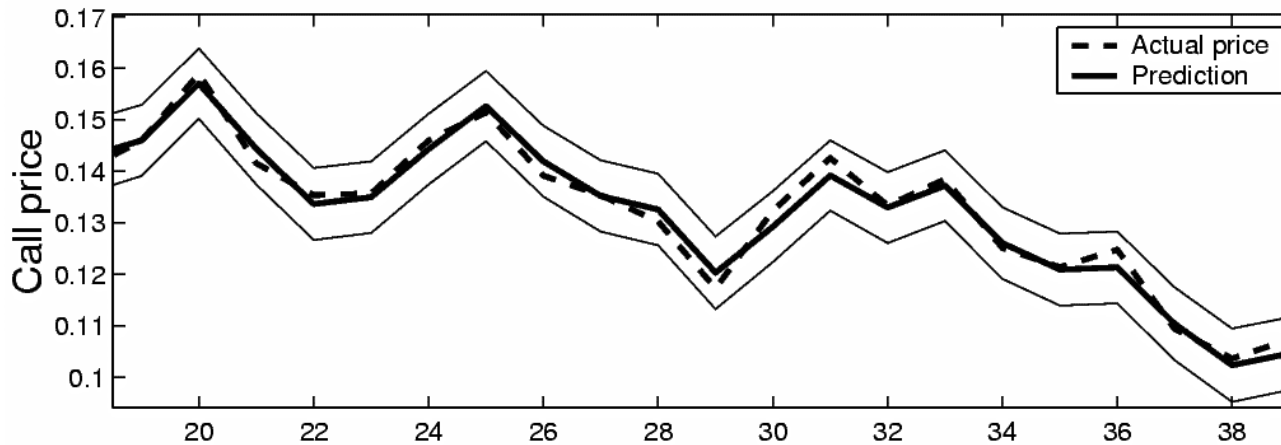
Experimental Results

- Options Pricing Experiment : (100 independent runs)

Option Type	Algorithm	NSE	
		mean	var
Call	Trivial	0.078	0.000
	Extended Kalman Filter (EKF)	0.037	0.000
	Unscented Kalman Filter (UKF)	0.037	0.000
	Particle Filter : generic	0.037	0.000
	Particle Filter : EKF proposal	0.092	0.508
	<i>Unscented Particle Filter</i>	<i>0.009</i>	<i>0.000</i>
Put	Trivial	0.035	0.000
	Extended Kalman Filter (EKF)	0.023	0.000
	Unscented Kalman Filter (UKF)	0.023	0.000
	Particle Filter : generic	0.023	0.000
	Particle Filter : EKF proposal	0.024	0.007
	<i>Unscented Particle Filter</i>	<i>0.008</i>	<i>0.000</i>

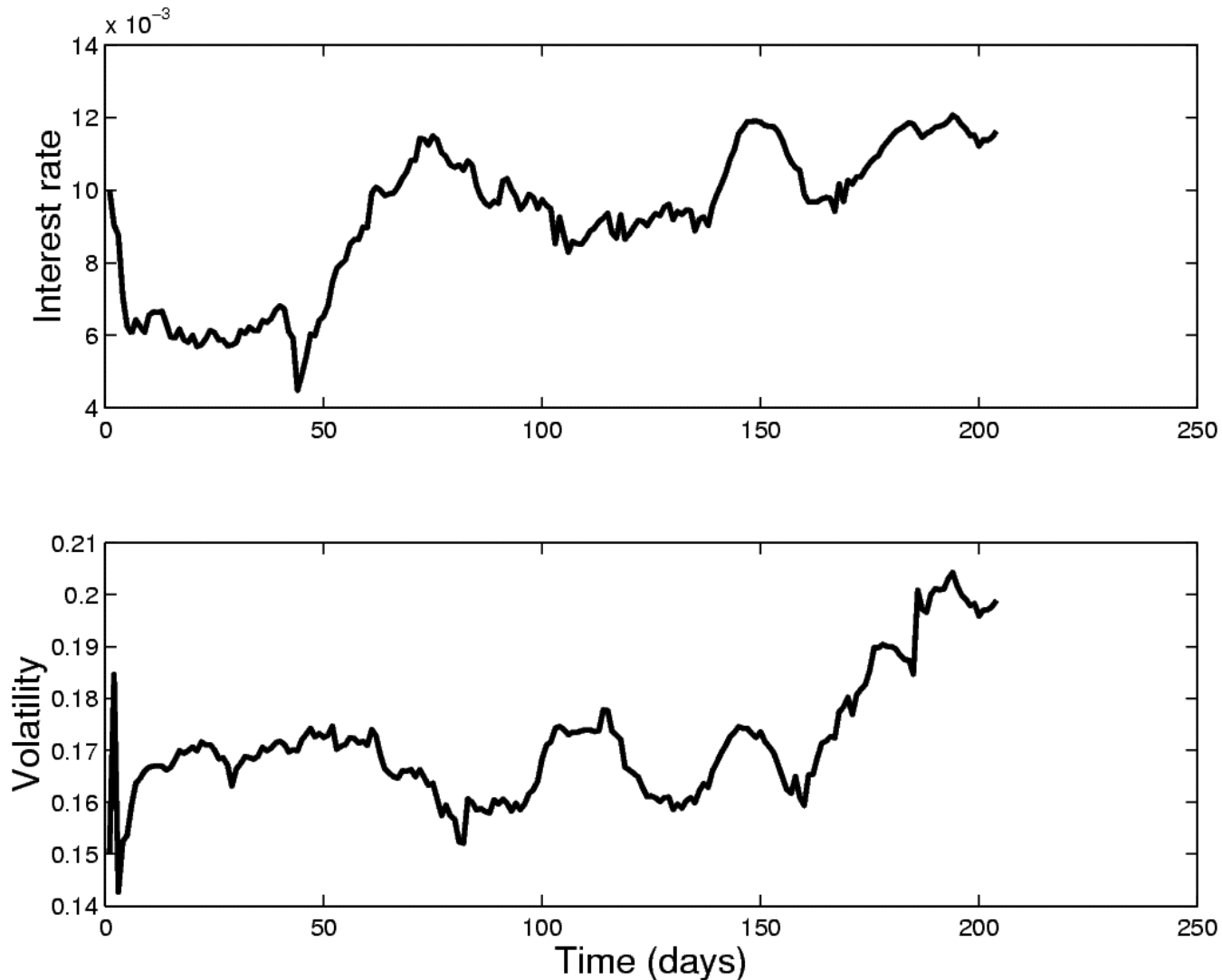
Experimental Results

- Options Pricing Experiment : UPF one-step-ahead predictions



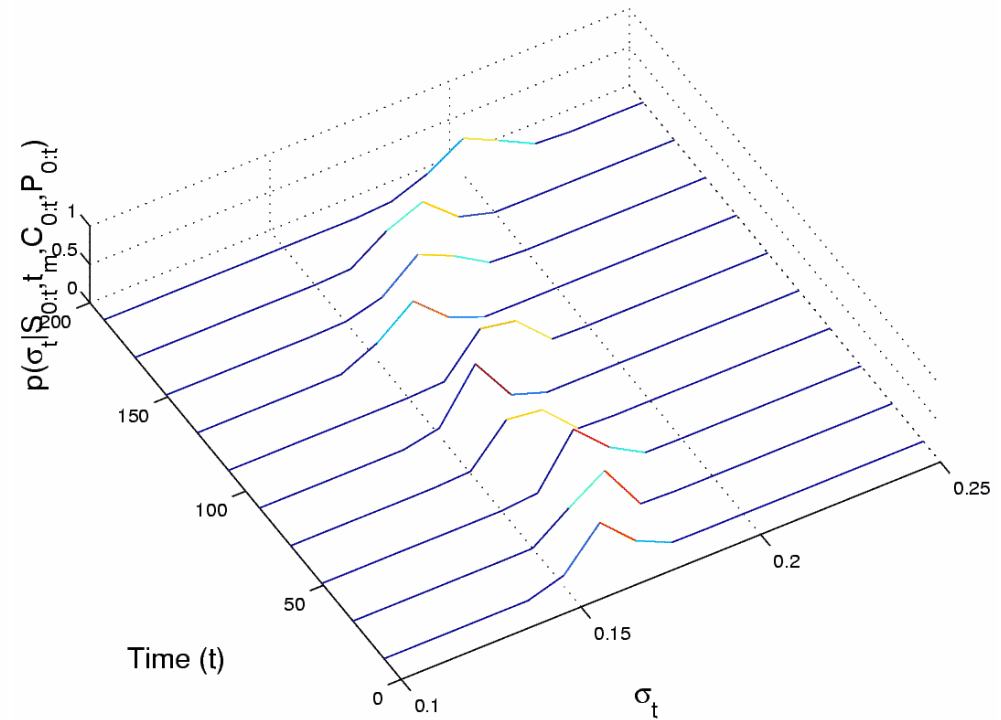
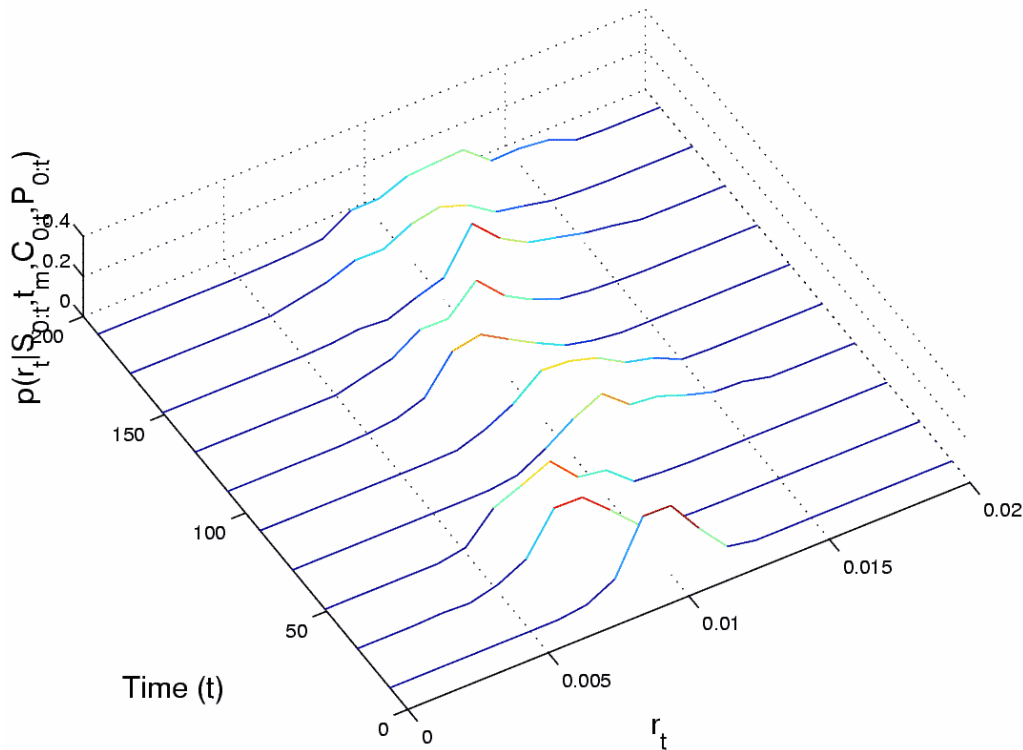
Experimental Results

- Options Pricing Experiment : Estimated interest rate and volatility



Experimental Results

- Options Pricing Experiment : Probability distributions of implied interest rate and volatility



Particle Filter Demos

- *Visual Dynamics Group, Oxford. (Andrew Blake)*

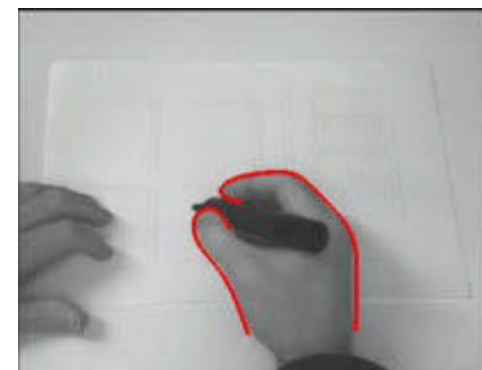
Tracking agile motion



Tracking motion against camouflage



Mixed state tracking



Conclusions

- Particle filters allow for a practical but complete representation of posterior probability distribution.
- Applicable to general nonlinear, non-Gaussian estimation problems where standard Gaussian approximations fail.
- Particle filters rely on importance sampling, so the proper choice of proposal distribution is very important:
 - ◆ Standard approach (i.e. transition prior proposal) fails when likelihood of new data is very peaked (accurate sensors) or for heavy-tailed noise sources.
 - ◆ EKF proposal : Incorporates new observations, but can diverge due to inaccurate and inconsistent state estimates.
 - ◆ *Unscented Particle Filter* : UKF proposal
 - More consistent and accurate state estimates.
 - Larger support overlap, can have heavier tailed distributions.
 - Theory predicts and experiments prove **significantly better** performance.

The End